Longest increasing subsequence as expectation of a simple nonlinear stochastic partial differential equation with a low noise intensity

E. Katzav,¹ S. Nechaev,^{2,*} and O. Vasilyev^{3,4}

¹Laboratoire de Physique Statistique de l'Ecole Normale Supérieure, CNRS UMR 8550, 24 rue Lhomond, 75231 Paris Cedex 05, France ²LPTMS, Université Paris Sud, 91405 Orsay Cedex, France

³Max-Planck-Institut für Metallforschung, Heisenbergstrasse 3, D-70569 Stuttgart, Germany

⁴Institut für Theoretische und Angewandte Physik, Universität Stuttgart, Pfaffenwaldring 57, D-70569 Stuttgart, Germany

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We report some observations concerning the statistics of longest increasing subsequences (LIS). We argue that the expectation of LIS, its variance, and apparently the full distribution function appears in statistical analysis of some simple nonlinear stochastic partial differential equation in the limit of very low noise intensity.

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I. INTRODUCTION

The last decade is marked by a breakthrough in the solution of a classical probabilistic problem, known as "Ulam problem" discussed in the mathematical literature over the years [1]. The general setting of the Ulam problem is as follows. Take the unit interval [0,1] and pick up from it one after another N random numbers $(N \ge 1)$ with uniform probability distribution. Having the sequence of N random numbers, one can extract from it the longest increasing subsequence (LIS) of k elements. The entries of this subsequence are not obliged to be the nearest neighbors. There are two basic questions of interest: (i) What is the expectation $\langle k(N) \rangle$ of LIS? (ii) What are the fluctuations of the mean length of LIS? The first question has been answered in 1975 by Vershik and Kerov [2] who have shown that $\langle k(N) \rangle = 2\sqrt{N}$ for $N \ge 1$. They derived this result by mapping the expectation of the LIS to the expectation of the first row of ensemble of Young tableaux over the Plancherel measure. If the initial random sequence contains repeated numbers, then the first line of the Young tableau corresponds to the longest nondecreasing subsequence.

The progress in treatment of fluctuations of LIS was the subject of recent investigations [3–8]. In these works an exact asymptotic form of a properly normalized full limiting distribution has been established (the so-called Tracy-Widom distribution) of the first row of the Young tableau over the Plancherel measure by mapping to the distribution of the largest eigenvalues of some classes of random matrices [3]. In particular it has been shown that the properly normalized variance of LIS has for $N \ge 1$ the asymptotic scaling form $\operatorname{var}[k(N)] \equiv \langle k^2(N) \rangle - \langle k(N) \rangle^2 = cN^{1/3}$, where $c = \langle [\chi - \langle \chi \rangle]^2 \rangle = 0.8132...$ and χ has the Tracy-Widom distribution for the Gaussian unitary ensemble.

We would like also to point out a deep connection between the distributions of edge states of random matrices and of statistical characteristics of some random growth models. In one of the most stimulating work [5], it has been shown that a (1+1)-dimensional model of directed polymers in random environment, which is in the Kardar-Parisi-Zhang (KPZ) universality class, has the Tracy-Widom distribution for the scaled height (energy). Around the same time the authors of [8] have found an exact mapping between a specific polynuclear growth (PNG) model and the LIS problem. The same Tracy-Widom distribution was reported in another class of (1+1)-dimensional growth models called "oriented digital boiling" model [9].

In our work we report some observations concerning the statistics of longest increasing subsequences. Namely, we provide an evidence that the expectation of LIS, its variance, and apparently the full distribution function appears in a statistical analysis of some properly scaled simple nonlinear stochastic partial differential equation (SPDE) in a limit of very low noise intensity.

The paper is organized as follows. In Sec. II we introduce all the necessary definitions. In particular, we describe LIS as the hight profile of some uniform discrete growth problem with nonlocal long-ranged interactions. In Sec. III we consider the discrete short-ranged asymmetric uniform growth process and derive the corresponding continuous space-time SPDE for the height profile. In Sec. IV we show that the properly scaled limit of an infinitely small noise in the derived SPDE adequately describes the growth with longranged interactions, i.e., the statistics of LIS. The discussion and conclusions are collected in Sec. V.

II. LIS AS A HEIGHT PROFILE IN A UNIFORM ASYMMETRIC NONLOCAL GROWTH PROCESS

The standard construction of a Young tableau for the set of real numbers is realized via the well known Robinson-Schensted-Knuth algorithm (RSK) which can be found in many textbook on representation theory, and for example, in [10]. The RSK algorithm ensures that the first row of the corresponding Young tableau would be the Longest Increasing Subsequence (LIS) for a given set of numbers. To be specific, consider the example of N=8 numbers: 4,4,3,6,7,3,4,2 taken at random with uniform probability dis-

^{*}Also at P.N. Lebedev Physical Institute of the Russian Academy of Sciences, 119991, Moscow, Russia.



FIG. 1. (a) Young tableau; (b) asymmetric ballistic deposition.

tribution from the support $\{1, \ldots, 9\}$. Following the RSK algorithm of the Young tableau construction, we get the Young tableau shown in Fig. 1(a). Let us describe now more geometrically obvious construction of LIS which refers to the discrete uniform asymmetric ballistic deposition of elementary cells with long-ranged interactions. This construction has appeared for the first time in [11].

Consider the (1+1)D model of ballistic deposition in which the columnar growth occurs sequentially on a linear substrate consisting of L columns with free boundary conditions. The time t is discrete and increases by 1 with every deposition event. We start with the flat initial condition, i.e., an empty substrate at t=0. At any stage of the growth, a column (say the column m) is chosen at random with probability p=1/L and a "cell" is deposited there which increases the height of this column by one unit: $h_m \rightarrow h_m + 1$. Once this "cell" is deposited, it screens all the sites at the same level in all the columns to its right from future deposition, i.e., the heights at all the columns to the right of the column m must be strictly greater than or equal to h_m+1 at all subsequent times. Formally such a growth is implemented by the following update rule. If the site m is chosen at time t for deposition, then

$$h_m(t+1) = \max\{h_m(t), h_{m-1}(t), \dots, h_1(t)\} + 1.$$
(1)

The model is anisotropic and long-ranged and evidently even the average height profile $\langle h_m(t) \rangle$ depends nontrivially on both the column number k and time t. Let us demonstrate now the a bijection between the longest nondecreasing subsequence in the sequence of N random numbers uniformly taken from the support $\{1, 2, 3, ..., L\}$ and the height in the uniformly growing heap with anisotropic infinite-ranged interactions in a bounding box containing L columns.

This bijection is defined by assigning the first line in the Young tableau to the "most top" (or "visible" from the lefthand side) blocks—see the Fig. 1(b). For the configuration shown in Fig. 1 we have the sequence of N=8 numbers: 4,4,3,6,7,3,4,2 taken at random from the set {1,...,9} (there are L=9 columns in the box). The "visible" blocks define the longest nondecreasing subsequence (LNS): 2,3,4,7. To be exact, if there are few LNSs, our construction extracts one of them—exactly as in the RSK scheme.



FIG. 2. Updating rules for the asymmetric NNN ballistic deposition defined in Eq. (2a).

As one can see in Fig. 1, the first line in the Young tableau (1) exactly matches the "most top" subsequence defined in our model and the distribution of the maximal height of a heap coincides with the distribution of the longest nondecreasing subsequence.

III. SHORT-RANGED ASYMMETRIC BALLISTIC GROWTH AND ITS CONTINUOUS LIMIT

Consider a one-dimensional discrete model of ballistic deposition with asymmetric (one-sided) next-nearest-neighboring (NNN) interactions

$$h_i(t+1) = \max\{h_i(t), h_{i-1}(t)\} + \eta_i(t),$$
 (2a)

where $h_j(t)$ is the height of a surface at the lattice position *j* and time *t* (the time is discrete as well) and $\eta_j(t)$ is a telegraphlike uncorrelated noise:

 $\overline{m(t)} = n$

$$\eta_j(t) = \begin{cases} 1 & \text{with probability } p, \\ 0 & \text{with probability } 1 - p, \end{cases}$$
(2b)

where

(

$$\eta_{j}(t) - p,$$

$$\overline{\eta_{j}(t_{1}) - \eta_{j}(t_{1})})(\eta_{m}(t_{2}) - \overline{\eta_{m}(t_{2})}) = p(1 - p)\delta_{j,m}\delta_{t_{1},t_{2}}.$$
(2c)

The bar denotes averaging over the distribution (2b). Equations (2a)-(2c) completely describe the updating rules for the NNN discrete ballistic deposition with the asymmetric interactions shown in Fig. 2.

In order to derive a continuous model corresponding to Eq. (2a), we represent the max{ \cdots } operator in Eq. (2a) at time *t* by a sgn function sgn[*z*]:

$$\max\{h_{j-1}, h_j\} = \frac{1}{2}h_{j-1}(1 + \operatorname{sgn}[h_{j-1} - h_j]) + \frac{1}{2}h_j(1 + \operatorname{sgn}[h_j - h_{j-1}]), \quad (3)$$

where

$$\operatorname{sgn}[z] = \begin{cases} +1 & \text{for } z > 0, \\ -1 & \text{for } z < 0. \end{cases}$$

Substituting Eq. (3) into Eq. (2a) we get after some simple algebra

$$\begin{aligned} h_{j}(t+1) - h_{j}(t) &= -\frac{1}{2} [h_{j}(t) - h_{j-1}(t)] \\ &+ \frac{1}{2} [h_{j}(t) - h_{j-1}(t)] \text{sgn}[h_{j}(t) - h_{j-1}(t)] \\ &+ \eta_{i}(t). \end{aligned} \tag{4}$$

In the last expression one can easily identify the finitedifference derivatives. Denoting the spatial and temporal increments by Δx and Δt and taking into account that sgn[az]=sgn[z] (for a>0), we arrive at the following expression

$$\partial_t h(x,t) = -\frac{1}{2} \frac{\Delta x}{\Delta t} \partial_x h(x,t) (1 - \operatorname{sgn}[\partial_x h(x,t)]) + \frac{1}{\Delta t} \eta(x,t),$$
(5)

where $\partial_t h(x,t)$ and $\partial_x h(x,t)$ denote the partial derivatives of h(x,t) with respect to t and x correspondingly. Setting in Eq. (5) $\Delta t = \Delta x = 1$ we arrive at the nonlinear stochastic partial differential equation (SPDE) which is a continuous analog of the discrete asymmetric ballistic deposition described by Eq. (2a). The noise in Eq. (5) is still defied by Eqs. (2b) and (2c).

It is known [12] that symmetric ballistic deposition model with NNN interactions in the continuous limit has some relevance to KPZ equation [13]. For better comparison of the KPZ equation with our SPDE one, rewrite Eq. (5) as follows:

$$\partial_t h(x,t) = \frac{1}{2} \left| \partial_x h(x,t) \right| - \frac{1}{2} \partial_x h(x,t) + \eta(x,t) \tag{6a}$$

or, equivalently,

$$\partial_t h(x,t) = \begin{cases} \eta(x,t) - \partial_x h(x,t) & \text{if } \partial_x h(x,t) < 0, \\ \eta(x,t) & \text{if } \partial_x h(x,t) > 0, \end{cases}$$
(6b)

with $\eta(x,t)$ given by Eqs. (2b) and (2c). One sees that the stochastic equation (6a) does not contain a diffusion term and is very different from the KPZ one. However, as will be shown below, the expectation $\langle h(x,t) \rangle$ of Eq. (6a) converges for $t \ge 1$ and $x \ge 1$ to the expectation of LIS for a noise $\eta(x,t)$ with very small intensity p ($p < L^{-2}$). The same is valid for the variance, var[h(x,t)], which demonstrates the KPZ scaling behavior for sufficiently large values of x. All that allows us to conjecture that the full distribution of h(x,t)for small *p* converges to a Tracy-Widom distribution of LIS. Let us note that the structure of Eq. (6b) "ideologically" resembles the structure of the one-dimensional Barenblatt model for the field u(x,t) with different diffusion constants for $\partial_t u(x,t) > 0$ and $\partial_t u(x,t) < 0$ [14]. However, in the Barenblatt model the diffusion coefficient experiences the break as a function of a partial derivative in t while in our model the partial derivative in x is responsible for the break. Since that difference we do not expect that the Barenbblatt model could have any signature of KPZ-type behavior.

IV. SPDE IN THE INFINITELY RARE NOISE REGIME AND LIS

In this section we show that the limit of a telegraphlike noise with small intensity in the stochastic partial differential equation (6a) for the "height" function h(x,t) has LIS statistics. To do that, it is convenient to describe LIS in terms of (1+1)D ballistic deposition (BD) set by Eq. (1).

Let us begin with the limiting case of absence of any noise in Eq. (6a). The noiseless equation (6a) is invariant under the scaling transformation $h(x,t) \rightarrow h(ax,at)$ for any *a*. Rewriting Eq. (6a) in a finite-difference form on the lattice, we have

$$h(x,t+\Delta t) = h(x,t) + \frac{1}{2}\Delta t \left| \frac{h(x,t) - h(x - \Delta x,t)}{\Delta x} \right|$$
$$-\frac{1}{2}\Delta t \left(\frac{h(x,t) - h(x - \Delta x,t)}{\Delta x} \right).$$
(7)

We can set $\Delta t = \Delta x = 1$ and x = j on the interval $0 \le j \le L$ without any loss of generality. The noiseless system defined by Eq. (7) has a characteristic time scale $\tau \sim L$ of reaching the equilibrium. Hence, after *t* time steps, where $t > \tau$, the system described by Eq. (7) reaches its stationary state generating a nondecreasing staircaselike profile from any initial state. This can be seen recursively:

$$h(j,t+1) = h(j-1,t) \quad \text{if } h(j-1,t) > h(j,t),$$

$$h(j,t+1) = h(j,t) \quad \text{if } h(j-1,t) < h(j,t). \tag{8}$$

Since Eq. (8) is valid for any $0 \le j \le L$, after $t \sim L$ time steps, if in the initial state h(j=L,t=0) < h(j=1,t=0) then the heights at x=L and x=1 equalize: h(j=L,t)=h(j=1,t). This signals the appearance of effective long-ranged spatial correlations in the system on characteristic time scales of $\tau \sim L$.

To show the equivalence of SPDE and LIS in a noisy regime with small intensity, it is convenient to introduce the enveloping functions $r_m(t)$ and r(x,t) for long-ranged (LR) BD and SPDE correspondingly:

$$r_m(t) = \max\{h_m(t), \dots, h_1(t)\} \text{ LR BD [Eq. (1)]},$$
$$r(x,t) = \max_{0 \le y \le x} \{h(y,t)\} \text{ SPDE.}$$
(9)

If the intensity, p, of the noise $\eta(x,t)$ is small, after each noise event the function h(x,t) has time to reach its relaxed form r(x,t) before the subsequent noise contribution is added. The corresponding intensity, p, can be easily estimated. Namely, the mean time interval, τ , between positive noise events $[\eta(x,t)=1]$ for a system of size L and noise intensity p is $\tau \sim 1/Lp$. In more details this question is discussed in Appendix A. Since the function h(x,t) reaches the relaxed state r(x,t) at $\tau > L$, we can estimate p as $p < L^{-2}$. Now Eqs. (1) and (6a) can be rewritten in the unified form:

$$h_m(t+1) = r_m(t) + 1$$
 long-ranged BD [Eq. (1)],

$$h(x,t + \Delta t) = r(x,t) + \eta(x,t)$$
 SPDE with $p < L^{-2}$. (10)

In this form the equivalence between long-ranged BD and SPDE in a low-intensity noise regime is clearly seen.

Thus, relying on the equivalence between statistics of LIS and low-intensity solutions of SPDE (6a), and knowing explicitly all the moments of the distribution of LIS (see, for example [5]), we get the following asymptotic expressions for the expectation, $\langle h(x,t) \rangle$, and for the variance, $\operatorname{var}[h(x,t)] \equiv \langle h(x,t)^2 \rangle - \langle h(x,t) \rangle^2$:

$$\langle h(x,t)\rangle = tp + 2\sqrt{xtp},\tag{11}$$

$$\operatorname{var}[h(x,t)] = c(xtp)^{1/3},$$
 (12)

where h(x,t) is the asymptotic solution of Eq. (6a) with the noise distribution (2c). In Eq. (12) $c = \langle [\chi - \langle \chi \rangle]^2 \rangle = 0.8132...$ and χ has the Tracy-Widom distribution for the Gaussian unitary ensemble. Note that for p=1/L and x=t=L ($L \ge 1$) we arrive in Eq. (11) at the Vershik-Kerov result for the expectation of LIS, $\langle h(L) \rangle = 2\sqrt{L}$. To be rigorous, this expression is beyond our approximation since we are limited by the noise intensity $p_{cr}=1/L^2$. So, it is more convenient to work with the general expressions (11) and (12).

We confirm our analytic predictions (11) and (12) by numerical computations, simulating directly the ballistic deposition process (1) and the stochastic partial differential equation (6a) with noise intensity p=0.1/L and initial condition h(x,t=0)=0. Note that the intensity of the noise in our numerical simulations is L/10 times larger than the upper estimate $p \sim 1/L^2$, however the agreement with theoretical results (11) and (12) is still very good. In particular, we plot in Fig. 3(a) the average profiles $\langle h(x,t) \rangle$ for long-ranged BD and SPDE as a function of x ($0 \le x \le L$) for $L = 10^3$ at time points $t=2 \times 10^4$, 6×10^4 , 10^5 . In the same figure we plot the analytic expression (11). In Figs. 3(b) and 3(c) we plot in double logarithmic scale the variance var[h(x,n)] $=\langle h(x,t)^2 \rangle - \langle h(x,t) \rangle^2$ as a function of x for the same time points t. It should be noted that due to the numerical procedure of the definition of the height, the numerical computation of the variance leads to the following interpolating expression for var[h(x,t)]= $a(t)+c(xpt)^{1/3}$. However, for $x \ge 1$ and t = const, we arrive asymptotically at the expression (12). That is why we compare the expression (12) with the numerical data for large x only [Fig. 3(c)] where the term a(t) is negligible. However, the agreement between long-ranged BD and SPDE is very good both for the expectation and the variance—see Figs. 3(a) and 3(b). All plots are dimensionless under the supposition $\Delta t = \Delta x = 1$.

V. CONCLUSION

We have argued by the sequence of mappings that the expectation and the variance of the random profile h(x,t), described by the stochastic nonlinear partial differential equation (6a) in the limit of a noise with very low intensity, coincides with the expectation (11) and the variance (12) of the longest nondecreasing subsequence of the sequence of random integers. This statement is also confirmed numerically. On the basis of the obtained results we conjecture that not only the first moments, but the full probability distribution function of the random variable h(x,t) coincides with the Tracy-Widom distribution appearing in largest eigenvalue statistics of ensembles of random matrices.

Let us note that our analysis of Eq. (6a) is a bit indirect. It would be very desirable to get expectation and variance (11)



FIG. 3. (a) The mean profile $\langle h(x,t) \rangle$ as a function of x for time points $t=2 \times 10^4$, 6×10^4 , 10^5 for BD and SPDE in comparison with the analytical results (11); (b), (c) The log-log plots of the variance var[h(x,t)] as a function of x at time points $t=2 \times 10^4$, 6×10^4 , 10^5 for BD and SPDE in comparison with the analytic result (12). Squares and dotted line correspond to the time t=20000 steps (\blacksquare , BD; \Box , SPDE). Circles and dashed line correspond to the time t=60000 steps (\blacklozenge , BD; \bigcirc , SPDE). Triangles and solid line correspond to the time t=100000 (\blacktriangle , BD; \triangle , SPDE). The length of the system is L=1000.

and (12) of the random variable h(x,t) directly from the solution of the Fokker-Planck equation which corresponds to the Langevin equation (6a). However on this way we have met some difficulties. Namely, rewrite Eq. (6a) as follows:

$$\partial_t h(x,t) = \frac{1}{2} \left| \partial_x h(x,t) \right| - \frac{1}{2} \partial_x h(x,t) + \xi(x,t) + p, \quad (13)$$

where $\xi(x,t) = \eta(x,t) - p$; $\xi(x,t) = 0$, $\xi(x,t)\xi(x',t') = 2D\delta_{x,x'}\delta_{t,t'}$, and $D = \frac{1}{2}p(1-p)$. Now we can write the formal expression for the Fokker-Planck variational equation for the function W(h,t) (see, for example [15]):

$$\frac{\partial W(h,t)}{\partial t} = -\int dx \frac{\delta}{\delta h} \left\{ \left(\frac{1}{2} |\partial_x h| - \frac{1}{2} \partial_x h + p \right) W(h,t) \right\} + D \int dx \frac{\delta^2}{\delta h^2} W(h,t).$$
(14)

One can further simplify this equation by transforming into \hat{W} defined as

$$\hat{W}(h,t) \equiv W(h,t)e^{-(p/D)\int h(x,t)dx},$$
(15)

and thus eliminating the p dependence from Eq. (14). This in turn can be rearranged in a form of two coupled equations:

$$\frac{\partial \hat{W}(h,t)}{\partial t} = D \int dx \frac{\delta^2 \hat{W}(h,t)}{\delta h^2} - \begin{cases} \int dx \partial_x h \frac{\delta \hat{W}(h,t)}{\delta h} & \text{if } \partial_x h < 0, \\ 0 & \text{if } \partial_x h > 0, \end{cases}$$
(16)

where we have taken into account that $(\delta/\delta h)[\partial_x h(x,t)]=0$. Although this equation takes now an apparently simple form, the inherent nonlinearity makes it extremely difficult to solve. When looking for stationary solutions [that satisfy $\partial \hat{W}(h,t)/\partial t=0$] it is easy to rule out the possibility of a solution that respects detailed balance. It is also easy to see that in the limit of small *D* a stationary solution (if exists) is likely to respect $\partial_x h > 0$. However, going beyond these simple observations has not been achieved so far.

The Fokker-Planck equation (16) corresponds to the stochastic process (6b) which can be visualized alternatively as a (2+1)-dimensional correlated growth. Using the finitedifference form (7) of Eq. (6b) on the lattice, where we set $\Delta x = \Delta t = 1$ and x = j, we arrive at the following stochastic recursion relation:

$$h(x,t+1) = h(x,t) + \eta(x,t) - \begin{cases} h(x,t) - h(x-1,t) & \text{if } h(x,t) < h(x-1,t), \\ 0 & \text{if } h(x,t) > h(x-1,t), \end{cases}$$
(17)

with the initial and boundary conditions h(x,t=0) = h(x=0,t)=0. Create now the initial configuration distributing the random variable $\eta(x,t)$ in the plane $x \ge 1, t \ge 1$. Since $\eta(x,t)$ takes the values 0 or 1, we can show $\eta=1$ by black unit segments as it is shown in Fig. 4. The initial configuration at t=1 is depicted in Fig. 4(a). Now applying Eq. (17) we can constrict recursively the configuration of the field h(x,t) at subsequent time moments. Few configurations of the field h(x,t) for given initial distribution of $\eta(x,t)$ is shown in Figs. 4(b)-4(d) for time moments t=2,3,4 and $x=1,\ldots,5$.

The (2+1)-dimensional stochastic growth shown in Fig. 4 bijectively corresponds to the short-ranged ballistic growth considered in Sec. III and, hence, in the limit of small intensity *p* should correspond to the statistics of LIS. It would be



FIG. 4. Visualization of the stochastic growth determined by the recursion relation (17) for x=1, ..., 5 and for several time slices (a) t=1, (b) t=2, (c) t=3, and (d) t=4. (Space and time are dimensionless, $\Delta x = \Delta t = 1$.)

very desirable to simulate numerically directly the growth model shown in Fig. 4 to check the validity of the conjectured bijection.

We complete our consideration with the model with intermediate length of one-sided "arm." The extension of the rules (2a) and (2b) are as follows:

$$h_j(t+1) = \max\{h_j(t), h_{j-1}(t), \dots, h_{j-k}(t)\} + \eta_j(t), \quad (18a)$$

where

$$\eta_j(t) = \begin{cases} 1 & \text{with probability } p, \\ 0 & \text{with probability } 1 - p. \end{cases}$$
(18b)

After some simple algebra reproduced in Appendix B we arrive at the following PDE:

$$\partial_t h(x,t) = \frac{k}{2} \left[\left| \partial_x h(x,t) \right| - \partial_x h(x,t) \right] + \eta(x,t).$$
(19)

Absorbing the parameter k/2 into the time scale, we see that we end up with the same equation as before [compare Eqs. (19) and (6a)] with rescaled noise. The expression (19) gives an additional support of our main result that the short-ranged and long-ranged models differ only with the time scale of the noise.

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APPENDIX A: EQUILIBRATION TIME FOR SPDE

In the limit of small p ($p \ll 1$) the actual mean profile $\langle h(x,t) \rangle$ tends to the curve $y(x,t)=tp+2\sqrt{xtp}$. However for larger values of p ($p \sim 1$) the resulting profile does not have

enough time to relax and hence is displaced below the curve y(x,t). The question which we address here concerns the estimation of the time interval, τ , between the subsequent deposition events sufficient to approach the stable solution y(x,t).

We can estimate the equilibration time, τ , from the following arguments. The slope of the mean profile in the stationary state is

$$\frac{\partial y(x,t)}{\partial x} = \sqrt{\frac{pt}{x}}.$$
 (A1)

Hence the average length of the horizonal "plateau" is about

$$l(x) \simeq \frac{1}{\frac{\partial y(x,t)}{\partial x}} \simeq \sqrt{\frac{x}{pt}}.$$
 (A2)

We define the equilibration time, τ , as a time *t* in Eq. (A2) during which the length, l(x), of the plateau becomes that of the order of the system size, *L*. Substituting for *x* in Eq. (A2) the maximal value x=L, we arrive at the following estimate:

$$L \simeq \sqrt{\frac{L}{p\tau}},$$

which gives us

$$\tau \simeq \frac{1}{Lp}.$$
 (A3)

Comparing the equilibration time (A3) with the characteristic time of the system relaxation, $\tau \sim L$, we arrive at the estimate for the critical noise intensity, p_{cr} ,

$$p_{\rm cr} \simeq \frac{1}{L^2}$$

below which the stochastic partial differential equation (6a) has long-ranged behavior typical for the LIS problem.

APPENDIX B: SPDE FOR GROWTH MODEL WITH INTERMEDIATE LENGTH OF THE "ARM"

Our starting point is Eq. (18a). Using a generalization of Eq. (3) we can write the max operator in the following form:

$$\max\{h_{j}, h_{j-1}, \dots, h_{j-k}\} = \frac{1}{2^{k}} \sum_{i=0}^{k} h_{j-i} \prod_{\substack{\ell=0\\\ell \neq i}}^{k} (1 + \operatorname{sgn}[h_{j-i} - h_{j-\ell}]).$$
(B1)

After some simple algebra, and using the fact that $sgn^{2}[z]=1$, we get

$$\sum_{i=0}^{k} h_{j-i} \prod_{\ell=0}^{k} (1 + \operatorname{sgn}[h_{j-i} - h_{j-\ell}])$$
$$= h_{j} \left(1 + \operatorname{sgn}\left[\frac{\partial h}{\partial x}\right]\right)^{k} + h_{j-k} \left(1 - \operatorname{sgn}\left[\frac{\partial h}{\partial x}\right]\right)^{k}. \quad (B2)$$

Using the identity

$$\left(1 \pm \operatorname{sgn}\left[\frac{\partial h}{\partial x}\right]\right)^{k} = 2^{k-1} \left(1 \pm \operatorname{sgn}\left[\frac{\partial h}{\partial x}\right]\right),$$
 (B3)

we arrive at the following expression:

$$h_j(t+1) - h_j(t) = \frac{1}{2} [h_j(t) - h_{j-k}(t)] \left(\text{sgn} \left\lfloor \frac{\partial h}{\partial x} \right\rfloor - 1 \right) + \eta_j(t),$$
(B4)

which becomes Eq. (19) in the continuum limit.

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